Statistical Data Mining and Machine Learning Hilary Term 2016

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Slides and other materials available at:

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Bayesian Learning

Maximum Likelihood Principle

• A generative model for training data $\mathcal{D} = \{(x_i, y_i)\}_{i=1}^n$ given a parameter vector θ :

$$y_i \sim (\pi_1, ..., \pi_K), \quad x|y_i \sim g_{y_i}(x) = p(x|\phi_{y_i})$$

- k-th class conditional density assumed to have a parametric form for $g_k(x) = p(x|\phi_k)$ and all parameters are given by $\theta = (\pi_1, \dots, \pi_K; \phi_1, \dots, \phi_K)$
- Generative process defines the **likelihood function**: the joint distribution of all the observed data $p(\mathcal{D}|\theta)$ given a parameter vector θ .
- Process of generative learning consists of computing the MLE $\widehat{\theta}$ of θ based on \mathcal{D} :

$$\widehat{\theta} = \operatorname*{argmax}_{\theta \in \Theta} p(\mathcal{D}|\theta)$$

We then use a plug-in approach to perform classification

$$f_{\widehat{\theta}}(x) = \underset{k \in \{1, \dots, K\}}{\operatorname{argmax}} \ \mathbb{P}_{\widehat{\theta}}(Y = k | X = x) = \underset{k \in \{1, \dots, K\}}{\operatorname{argmax}} \ \frac{\widehat{\pi}_k p(x | \widehat{\phi}_k)}{\sum_{j=1}^K \widehat{\pi}_j p(x | \widehat{\phi}_j)}$$

The Bayesian Learning Framework

- Being Bayesian: treat parameter vector θ as a random variable: process of learning is then computation of the posterior distribution $p(\theta|\mathcal{D})$.
- In addition to the likelihood $p(\mathcal{D}|\theta)$ need to specify a **prior distribution** $p(\theta)$.
- Posterior distribution is then given by the Bayes Theorem:

$$p(\theta|\mathcal{D}) = \frac{p(\mathcal{D}|\theta)p(\theta)}{p(\mathcal{D})}$$

• Likelihood: $p(\mathcal{D}|\theta)$

• Posterior: $p(\theta|\mathcal{D})$

• Prior: $p(\theta)$

- Marginal likelihood: $p(\mathcal{D}) = \int_{\Theta} p(\mathcal{D}|\theta) p(\theta) d\theta$
- Summarizing the posterior:
 - Posterior mode: $\widehat{\theta}^{MAP} = \operatorname{argmax}_{\theta \in \Theta} p(\theta | \mathcal{D})$ (maximum a posteriori).
 - Posterior mean: $\widehat{\theta}^{\text{mean}} = \mathbb{E} [\theta | \mathcal{D}].$
 - Posterior variance: $Var[\theta|\mathcal{D}]$.

- A simple example: We have a coin with probability ϕ of coming up heads. Model coin tosses as i.i.d. Bernoullis, 1 = head, 0 = tail.
- Estimate ϕ given a dataset $\mathcal{D} = \{x_i\}_{i=1}^n$ of tosses.

$$p(\mathcal{D}|\phi) = \phi^{n_1}(1-\phi)^{n_0}$$

with $n_i = \sum_{i=1}^n \mathbb{1}(x_i = j)$.

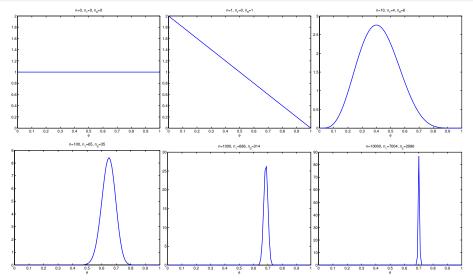
Maximum Likelihood estimate:

$$\hat{\phi}^{\mathsf{ML}} = \frac{n_1}{n}$$

• Bayesian approach: treat the unknown parameter ϕ as a random variable. Simple prior: $\phi \sim \mathsf{Uniform}[0,1]$, i.e., $p(\phi) = 1$ for $\phi \in [0,1]$. Posterior distribution:

$$p(\phi|\mathcal{D}) = \frac{p(\mathcal{D}|\theta)p(\theta)}{p(\mathcal{D})} = \frac{\phi^{n_1}(1-\phi)^{n_0} \cdot 1}{p(\mathcal{D})}, \quad p(\mathcal{D}) = \int_0^1 \phi^{n_1}(1-\phi)^{n_0}d\phi = \frac{(n+1)!}{n_1!n_0!}$$

Posterior is a Beta $(n_1 + 1, n_0 + 1)$ distribution: $\widehat{\phi}^{\text{mean}} = \frac{n_1 + 1}{n + 2}$.



Posterior becomes behaves like the ML estimate as dataset grows and is peaked at true value $\phi^* = .7$.

- All Bayesian reasoning is based on the posterior distribution.
 - Posterior mode: $\widehat{\phi}^{MAP} = \frac{n_1}{n_2}$
 - Posterior mean: $\widehat{\phi}^{\text{mean}} = \frac{n_1+1}{n+2}$
 - Posterior variance: $Var[\phi|\mathcal{D}] = \frac{1}{n+3} \widehat{\phi}^{mean} (1 \widehat{\phi}^{mean})$
 - $(1-\alpha)$ -credible regions: $(l,r) \subset [0,1]$ s.t. $\int_{l}^{r} p(\theta|\mathcal{D}) d\theta = 1-\alpha$.
- Consistency: Assuming that the true parameter value ϕ^* is given a non-zero density under the prior, the posterior distribution concentrates around the true value as $n \to \infty$.
- Rate of convergence?

• The **posterior predictive distribution** is the conditional distribution of x_{n+1} given $\mathcal{D} = \{x_i\}_{i=1}^n$:

$$p(x_{n+1}|\mathcal{D}) = \int_0^1 p(x_{n+1}|\phi, \mathcal{D})p(\phi|\mathcal{D})d\phi$$
$$= \int_0^1 p(x_{n+1}|\phi)p(\phi|\mathcal{D})d\phi$$
$$= (\widehat{\phi}^{\mathsf{mean}})^{x_{n+1}}(1 - \widehat{\phi}^{\mathsf{mean}})^{1 - x_{n+1}}$$

• We predict on new data by **averaging** the predictive distribution over the posterior. Accounts for uncertainty about ϕ .

- In this example, the posterior distribution has a known analytic form and is in the same Beta family as the prior: Uniform[0, 1] ≡ Beta(1, 1).
- An example of a conjugate prior.
- A Beta distribution $\mathrm{Beta}(a,b)$ with parameters a,b>0 is an exponential family distribution with density

$$p(\phi|a,b) = \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \phi^{a-1} (1-\phi)^{b-1}$$

where $\Gamma(t) = \int_0^\infty u^{t-1} e^{-u} du$ is the gamma function.

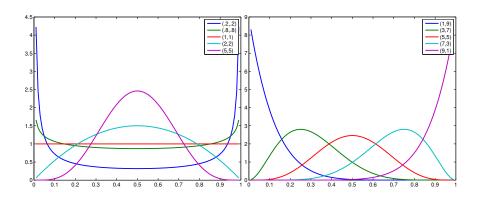
• If the prior is $\phi \sim \text{Beta}(a, b)$, then the posterior distribution is

$$p(\phi|\mathcal{D}, a, b) = \propto \phi^{a+n_1-1} (1-\phi)^{b+n_0-1}$$

so is Beta($a + n_1, b + n_0$).

• Hyperparameters a and b are **pseudo-counts**, an imaginary initial sample that reflects our prior beliefs about ϕ .

Beta Distributions



Bayesian Inference on the Categorical Distribution

• Suppose we observe $\mathcal{D} = \{y_i\}_{i=1}^n$ with $y_i \in \{1, \dots, K\}$, and model them as i.i.d. with pmf $\pi = (\pi_1, \dots, \pi_K)$:

$$p(\mathcal{D}|\pi) = \prod_{i=1}^{n} \pi_{y_i} = \prod_{k=1}^{K} \pi_k^{n_k}$$

with $n_k = \sum_{i=1}^n \mathbb{1}(y_i = k)$ and $\pi_k > 0$, $\sum_{k=1}^K \pi_k = 1$.

• The conjugate prior on π is the Dirichlet distribution $\mathrm{Dir}(\alpha_1,\ldots,\alpha_K)$ with parameters $\alpha_k>0$, and density

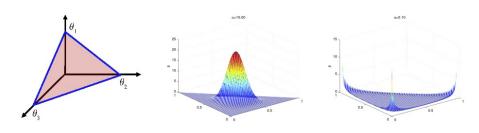
$$p(\pi) = \frac{\Gamma(\sum_{k=1}^{K} \alpha_k)}{\prod_{k=1}^{K} \Gamma(\alpha_k)} \prod_{k=1}^{K} \pi_k^{\alpha_k - 1}$$

on the probability simplex $\{\pi : \pi_k > 0, \sum_{k=1}^K \pi_k = 1\}$.

- The posterior is also Dirichlet $Dir(\alpha_1 + n_1, \dots, \alpha_K + n_K)$.
- Posterior mean is

$$\widehat{\pi}_k^{\mathsf{mean}} = \frac{\alpha_k + n_k}{\sum_{i=1}^K \alpha_i + n_i}.$$

Dirichlet Distributions



- (A) Support of the Dirichlet density for K = 3.
- (B) Dirichlet density for $\alpha_k = 10$.
- (C) Dirichlet density for $\alpha_k = 0.1$.

Naïve Bayes

Return to the spam classification example with two-class naïve Bayes

$$p(x_i|\phi_k) = \prod_{i=1}^p \phi_{kj}^{x_i^{(j)}} (1 - \phi_{kj})^{1 - x_i^{(j)}}.$$

• Set $n_k = \sum_{i=1}^n \mathbb{1}\{y_i = k\}, n_{kj} = \sum_{i=1}^n \mathbb{1}(y_i = k, x_i^{(j)} = 1)$. MLE is:

$$\hat{\pi}_k = \frac{n_k}{n},$$
 $\hat{\phi}_{kj} = \frac{\sum_{i:y_i=k} x_i^{(j)}}{n_k} = \frac{n_{kj}}{n_k}.$

• One problem: if the ℓ -th word did not appear in documents labelled as class k then $\hat{\phi}_{k\ell}=0$ and

$$\mathbb{P}(Y = k | X = x \text{ with } \ell\text{-th entry equal to 1})$$

$$\propto \hat{\pi}_k \prod_{i=1}^p \left(\hat{\phi}_{kj}\right)^{x^{(i)}} \left(1 - \hat{\phi}_{kj}\right)^{1 - x^{(i)}} = 0$$

i.e. we will never attribute a new document containing word ℓ to class k (regardless of other words in it).

Bayesian Inference on Naïve Bayes model

• Under the Naïve Bayes model, the joint distribution of labels $y_i \in \{1, ..., K\}$ and data vectors $x_i \in \{0, 1\}^p$ is

$$\prod_{i=1}^{n} p(x_i, y_i) = \prod_{i=1}^{n} \prod_{k=1}^{K} \left(\pi_k \prod_{j=1}^{p} \phi_{kj}^{x_i^{(j)}} (1 - \phi_{kj})^{1 - x_i^{(j)}} \right)^{\mathbb{1}(y_i = k)}$$

$$= \prod_{k=1}^{K} \pi_k^{n_k} \prod_{j=1}^{p} \phi_{kj}^{n_{kj}} (1 - \phi_{kj})^{n_k - n_{kj}}$$

where
$$n_k = \sum_{i=1}^n \mathbb{1}(y_i = k)$$
, $n_{kj} = \sum_{i=1}^n \mathbb{1}(y_i = k, x_i^{(j)} = 1)$.

- For conjugate prior, we can use $\mathrm{Dir}((\alpha_k)_{k=1}^K)$ for π , and $\mathrm{Beta}(a,b)$ for ϕ_{kj} independently.
- Because the likelihood factorizes, the posterior distribution over π and (ϕ_{kj}) also factorizes, and posterior for π is $\mathrm{Dir}((\alpha_k+n_k)_{k=1}^K)$, and for ϕ_{kj} is $\mathrm{Beta}(a+n_{kj},b+n_k-n_{kj})$.

Bayesian Inference on Naïve Bayes model

• Given $\mathcal{D} = \{(x_i, y_i)\}_{i=1}^n$, want to predict a label \tilde{y} for a new document \tilde{x} . We can calculate

$$p(\tilde{x}, \tilde{y} = k|\mathcal{D}) = p(\tilde{y} = k|\mathcal{D})p(\tilde{x}|\tilde{y} = k, \mathcal{D})$$

with

$$p(\tilde{y} = k | \mathcal{D}) = \frac{\alpha_k + n_k}{\sum_{l=1}^K \alpha_l + n}$$
$$p(\tilde{x}^{(j)} = 1 | \tilde{y} = k, \mathcal{D}) = \frac{a + n_{kj}}{a + b + n_k}$$

Predicted class is

$$p(\tilde{y} = k | \tilde{x}, \mathcal{D}) = \frac{p(\tilde{y} = k | \mathcal{D})p(\tilde{x} | \tilde{y} = k, \mathcal{D})}{p(\tilde{x} | \mathcal{D})}$$

 Compared to ML plug-in estimator, pseudocounts help to "regularize" probabilities away from extreme values.

Bayesian Learning and Regularization

• Consider a Bayesian approach to logistic regression: introduce a multivariate normal prior for weight vector $w \in \mathbb{R}^p$, and a uniform (improper) prior for offset $b \in \mathbb{R}$. The prior density is:

$$p(b, w) = 1 \cdot (2\pi\sigma^2)^{-\frac{\rho}{2}} \exp\left(-\frac{1}{2\sigma^2} \|w\|_2^2\right)$$

The posterior is

$$p(b, w|\mathcal{D}) \propto \exp\left(-\frac{1}{2\sigma^2}||w||_2^2 - \sum_{i=1}^n \log(1 + \exp(-y_i(b + w^{\top}x_i)))\right)$$

- The posterior mode is equivalent to minimizing the L₂-regularized empirical risk.
- Regularized empirical risk minimization is (often) equivalent to having a prior and finding a MAP estimate of the parameters.
 - L₂ regularization multivariate normal prior.
 - L₁ regularization multivariate Laplace prior.
- From a Bayesian perspective, the MAP parameters are just one way to summarize the posterior distribution.

Bayesian Model Selection

- A model \mathcal{M} with a given set of parameters $\theta_{\mathcal{M}}$ consists of both the likelihood $p(\mathcal{D}|\theta_{\mathcal{M}})$ and the prior distribution $p(\theta_{\mathcal{M}})$.
 - One example model would consist of all Gaussian mixtures with K components and equal covariance (LDA): $\theta_{\text{LDA}} = (\pi_1, \dots, \pi_K; \mu_1, \dots, \mu_K; \Sigma)$, along with a prior on θ ; another would allow different covariances (QDA) $\theta_{\text{QDA}} = (\pi_1, \dots, \pi_K; \mu_1, \dots, \mu_K; \Sigma_1, \dots, \Sigma_K)$.
- The posterior distribution

$$p(\theta_{\mathcal{M}}|\mathcal{D}, \mathcal{M}) = \frac{p(\mathcal{D}|\theta_{\mathcal{M}}, \mathcal{M})p(\theta_{\mathcal{M}}|\mathcal{M})}{p(\mathcal{D}|\mathcal{M})}$$

• Marginal probability of the data under \mathcal{M} (Bayesian model evidence):

$$p(\mathcal{D}|\mathcal{M}) = \int_{\Theta} p(\mathcal{D}|\theta_{\mathcal{M}}, \mathcal{M}) p(\theta_{\mathcal{M}}|\mathcal{M}) d\theta$$

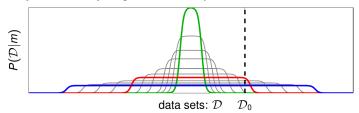
Compare models using their **Bayes factors** $\frac{p(\mathcal{D}|\mathcal{M})}{p(\mathcal{D}|\mathcal{M}')}$

Bayesian Occam's Razor

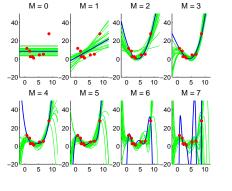
 Occam's Razor: of two explanations adequate to explain the same set of observations, the simpler should be preferred.

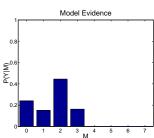
$$p(\mathcal{D}|\mathcal{M}) = \int_{\Theta} p(\mathcal{D}|\theta_{\mathcal{M}}, \mathcal{M}) p(\theta_{\mathcal{M}}|\mathcal{M}) d\theta$$

- Model evidence $p(\mathcal{D}|\mathcal{M})$ is the probability that a set of randomly selected parameter values inside the model would generate dataset \mathcal{D} .
- Models that are too simple are unlikely to generate the observed dataset.
- Models that are too complex can generate many possible dataset, so again, they are unlikely to generate that particular dataset at random.



Bayesian model comparison: Occam's razor at work





Bayesian Learning - Discussion

- Use probability distributions to reason about uncertainties of parameters (latent variables and parameters are treated in the same way).
- Model consists of the likelihood function and the prior distribution on parameters: allows to integrate prior beliefs and domain knowledge.
- Bayesian computation most posteriors are intractable, and posterior needs to be approximated by:
 - Monte Carlo methods (MCMC and SMC).
 - Variational methods (variational Bayes, belief propagation etc).
- Prior usually has hyperparameters, i.e., $p(\theta) = p(\theta|\psi)$. How to choose ψ ?
 - Be Bayesian about ψ as well choose a hyperprior $p(\psi)$ and compute $p(\psi|\mathcal{D})$.
 - $\bullet \ \ \text{Maximum Likelihood II} \hat{\psi} = \text{argmax}_{\psi \in \Psi} p(\mathcal{D}|\psi).$

$$p(\mathcal{D}|\psi) = \int p(\mathcal{D}|\theta)p(\theta|\psi)d\theta$$
$$p(\psi|\mathcal{D}) = \frac{p(\mathcal{D}|\psi)p(\psi)}{p(\mathcal{D})}$$

Bayesian Learning – Further Reading

- Videolectures by Zoubin Ghahramani:
 Bayesian Learning and Graphical models.
- Gelman et al. Bayesian Data Analysis.
- Kevin Murphy. Machine Learning: a Probabilistic Perspective.
- E. T. Jaynes. Probability Theory: The Logic of Science.

Parametric vs Nonparametric models

• **Parametric models** have a fixed finite number of parameters, regardless of the dataset size. In the Bayesian setting, given the parameter vector θ , the predictions are independent of the data \mathcal{D} .

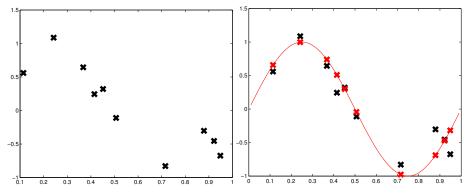
$$p(\tilde{x}, \theta|\mathcal{D}) = p(\theta|\mathcal{D})p(\tilde{x}|\theta)$$

Parameters can be thought of as a data summary: communication channel flows from data to the predictions through the parameters. Model-based learning (e.g., mixture of K multivariate normals)

 Nonparametric models allow the number of "parameters" to grow with the dataset size. Alternatively, predictions depend on the data (and the hyperparameters).

Memory-based learning (e.g., kernel density estimation)

Regression



- We are given a dataset $\mathcal{D} = \{(x_i, y_i)\}_{i=1}^n, x_i \in \mathbb{R}^p, y_i \in \mathbb{R}.$
- Regression: learn the underlying real-valued function f(x).

Different Flavours of Regression

 We can model response y_i as a noisy version of the underlying function f evaluated at input x_i:

$$y_i|f,x_i \sim \mathcal{N}(f(x_i),\sigma^2)$$

Appropriate loss: $L(y, f(x)) = (y - f(x))^2$

- Frequentist Parametric approach: model f as f_{θ} for some parameter vector θ . Fit θ by ML / ERM with squared loss (linear regression).
- Frequentist Nonparametric approach: model f as the unknown parameter taking values in an infinite-dimensional space of functions. Fit f by regularized ML / ERM with squared loss (kernel ridge regression)
- Bayesian Parametric approach: model f as f_{θ} for some parameter vector θ . Put a prior on θ and compute a posterior $p(\theta|\mathcal{D})$ (Bayesian linear regression).
- Bayesian Nonparametric approach: treat f as the random variable taking values in an infinite-dimensional space of functions. Put a prior over functions $f \in \mathcal{F}$, and compute a posterior $p(f|\mathcal{D})$ (Gaussian Process regression).

- Just work with the function values at the inputs $\mathbf{f} = (f(x_1), \dots, f(x_n))^{\top}$
- What properties of the function can we incorporate?
 - Multivariate normal prior on f:

$$\mathbf{f} \sim \mathcal{N}(0, \mathbf{K})$$

Use a kernel function k to define K:

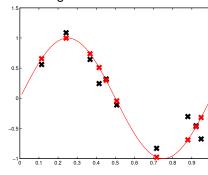
$$\mathbf{K}_{ij} = k(x_i, x_j)$$

• Expect regression functions to be smooth: If x and x' are close by, then f(x) and f(x') have similar values, i.e. strongly correlated.

$$\begin{pmatrix} f(x) \\ f(x') \end{pmatrix} \sim \mathcal{N} \left(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} k(x,x) & k(x,x') \\ k(x',x) & k(x',x') \end{pmatrix} \right)$$

In particular, want $k(x, x') \approx k(x, x) = k(x', x')$.

The prior $p(\mathbf{f})$ encodes our prior knowledge about the function.



Model:

$$\mathbf{f} \sim \mathcal{N}(0, \mathbf{K})$$

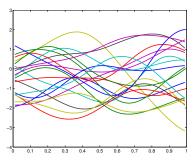
 $y_i | f_i \sim \mathcal{N}(f_i, \sigma^2)$

- What does a multivariate normal prior mean?
- Imagine x forms an infinitesimally dense grid of data space. Simulate prior draws

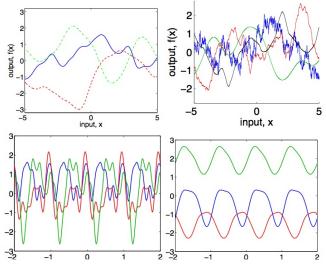
$$\boldsymbol{f} \sim \mathcal{N}(0,\boldsymbol{K})$$

Plot f_i vs x_i for i = 1, ..., n.

 The corresponding prior over functions is called a Gaussian Process (GP): any finite number of evaluations of which follow a Gaussian distribution.



Different kernels lead to different function characteristics.



Carl Rasmussen. Tutorial on Gaussian Processes at NIPS 2006.

$$\mathbf{f}|\mathbf{x} \sim \mathcal{N}(0, \mathbf{K})$$

 $\mathbf{y}|\mathbf{f} \sim \mathcal{N}(\mathbf{f}, \sigma^2 I)$

Posterior distribution:

$$\mathbf{f}|\mathbf{y} \sim \mathcal{N}(\mathbf{K}(\mathbf{K} + \sigma^2 I)^{-1}\mathbf{y}, \mathbf{K} - \mathbf{K}(\mathbf{K} + \sigma^2 I)^{-1}\mathbf{K})$$

 Posterior predictive distribution: Suppose x' is a test set. We can extend our model to include the function values f' at the test set:

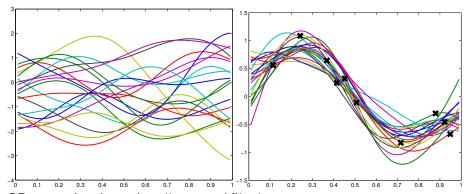
$$\begin{pmatrix} \mathbf{f} \\ \mathbf{f'} \end{pmatrix} | \mathbf{x}, \mathbf{x'} \sim \mathcal{N} \begin{pmatrix} \begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} \mathbf{K}_{\mathbf{x}\mathbf{x}} & \mathbf{K}_{\mathbf{x}\mathbf{x'}} \\ \mathbf{K}_{\mathbf{x'}\mathbf{x}} & \mathbf{K}_{\mathbf{x'}\mathbf{x'}} \end{pmatrix} \end{pmatrix}$$

$$\mathbf{y} | \mathbf{f} \sim \mathcal{N}(\mathbf{f}, \sigma^2 I)$$

where $\mathbf{K}_{\mathbf{x}\mathbf{x}'}$ is matrix with (i,j)-th entry $k(x_i,x_i')$.

Some manipulation of multivariate normals gives:

$$\mathbf{f'}|\mathbf{y} \sim \mathcal{N}\left(\mathbf{K_{x'x}}(\mathbf{K_{xx}} + \sigma^2 I)^{-1}\mathbf{y}, \mathbf{K_{x'x'}} - \mathbf{K_{x'x}}(\mathbf{K_{xx}} + \sigma^2 I)^{-1}\mathbf{K_{xx'}}\right)$$



GP regression demo: http://www.tmpl.fi/gp/

- A whirlwind journey through data mining and machine learning techniques:
 - Unsupervised learning: PCA, MDS, Isomap, Hierarchical clustering, K-means, mixture modelling, EM algorithm.
 - Supervised learning: LDA, QDA, naïve Bayes, logistic regression, SVMs, kernel methods, kNN, deep neural networks, Gaussian processes, decision trees, ensemble methods: random forests, bagging, stacking, dropout and boosting.
 - Conceptual frameworks: prediction, performance evaluation, generalization, overfitting, regularization, model complexity, validation and cross-validation, bias-variance tradeoff.
 - Theory: decision theory, statistical learning theory, convex optimization,
 Bayesian vs. frequentist learning, parametric vs non-parametric learning.

• Further resources:

- Machine Learning Summer Schools, videolectures.net.
- Conferences: NIPS, ICML, UAI, AISTATS.
- Mailing list: ml-news.

Thank You!